

Graphical Models

Gaussian Network Models

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Learning objectives

- multivariate Gaussian density:
 - different parametrizations
 - marginalization and conditioning
 - expression as Markov & Bayesian networks

Univariate Gaussian density

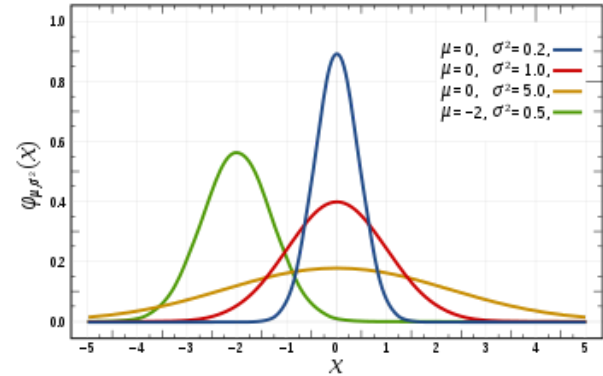
$$p(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

- motivated by central limit theorem
- max-entropy dist. with a fixed variance

$$\mu \in \mathfrak{R}, \sigma^2 > 0$$

$$\mathbb{E}[X] = \mu,$$


$$\mathbb{E}[X^2] - \mathbb{E}[X]^2 = \sigma^2$$



Multivariate Gaussian

$\mathbf{x} \in \mathfrak{R}^n$ is a column vector (*convention*)

$$p(\mathbf{x}; \mu, \Sigma) = \frac{1}{\sqrt{|2\pi\Sigma|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mu)^T \Sigma^{-1}(\mathbf{x} - \mu)\right)$$


$$(2\pi)^{-\frac{n}{2}} |\Sigma|^{-\frac{1}{2}}$$

compre to $p(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$

Multivariate Gaussian; **sufficient statistics**

$$p(\mathbf{x}; \mu, \Sigma) = \frac{1}{\sqrt{|2\pi\Sigma|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mu)^T \Sigma^{-1}(\mathbf{x} - \mu)\right)$$

$$\mu = \mathbb{E}[\mathbf{X}]$$

$$\Sigma = \underbrace{\mathbb{E}[\mathbf{X}\mathbf{X}^T]}_{n \times n} - \underbrace{\mathbb{E}[\mathbf{X}]\mathbb{E}[\mathbf{X}^T]}_{n \times n} \quad \text{the covariance matrix} \quad \left| \begin{array}{l} \Sigma_{i,i} = \text{Var}(X_i) \\ \Sigma_{i,j} = \text{Cov}(X_i, X_j) \end{array} \right.$$

only captures these two statistics

Multivariate Gaussian; **covariance matrix**

since $\mathbf{y}^T \Sigma \mathbf{y} = (\mathbf{y}^T \mathbb{E}[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])^T] \mathbf{y}) = a^2 > 0$

↑
move this expectation out

Multivariate Gaussian; **covariance matrix**

since $\mathbf{y}^T \Sigma \mathbf{y} = (\mathbf{y}^T \mathbb{E}[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])^T] \mathbf{y}) = a^2 > 0$
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move this expectation out

- $\Sigma \succ 0$
- is **symmetric positive definite** (PD) $\mathbf{y}^T \Sigma \mathbf{y} > 0 \quad \forall \mathbf{y}; \|\mathbf{y}\| > 0$
 - the inverse of a PD matrix is PD
 - the **precision matrix** $\Lambda = \Sigma^{-1} \succ 0$
 - is diagonalized by orthonormal matrices

Multivariate Gaussian; covariance matrix

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 - is diagonalized by orthonormal matrices

$$\Sigma = \mathbf{Q} \mathbf{D} \mathbf{Q}^T$$

↓ ↓
 diagonal

- orthogonal rows & columns of unit norm $\mathbf{Q} \mathbf{Q}^T = \mathbf{Q}^T \mathbf{Q} = \mathbf{I}$
- rotation and reflection

Multivariate Gaussian; covariance matrix

$$\Sigma = Q D Q^T$$

↓ diagonal (scaling)

- orthogonal rows & columns of unit norm $Q Q^T = Q^T Q = I$
- rotation and reflection

Scaling along axes in some rotated/reflected coordinate system

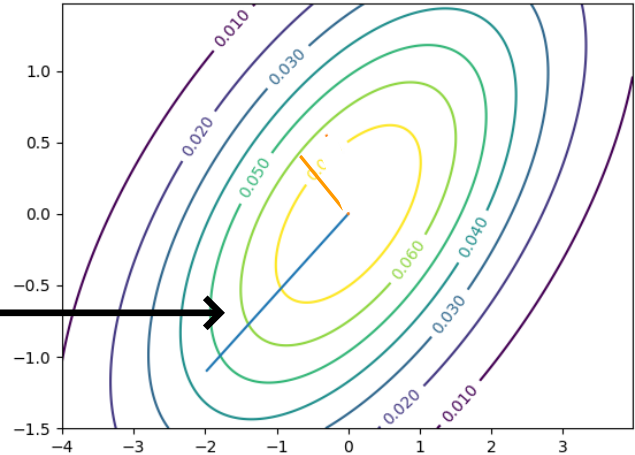
Multivariate Gaussian; **example**

$$p(\mathbf{x}; \mu, \Sigma) = \frac{1}{\sqrt{|2\pi\Sigma|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mu)^T \Sigma^{-1}(\mathbf{x} - \mu)\right)$$

$$\Sigma = \begin{bmatrix} 4, & 2 \\ 2, & \frac{1}{2} \end{bmatrix} \approx \begin{bmatrix} -.87, & -.48 \\ -.48, & .87 \end{bmatrix} \begin{bmatrix} 5.1, & 0 \\ 0, & .39 \end{bmatrix} \begin{bmatrix} -.87, & -.48 \\ -.48, & .87 \end{bmatrix}^T$$

$Q \qquad D \qquad Q^T$

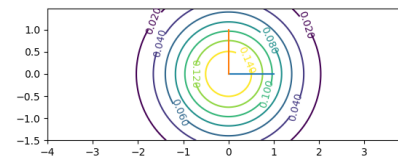
$$\sqrt{5.1} \times [-.87, .48]$$



Multivariate Gaussian; **from univariates**

given n univariate Gaussians

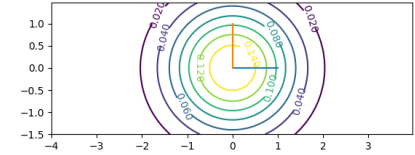
$$\mathbf{X} \sim \mathcal{N}(\mathbf{0}, I)$$



Multivariate Gaussian; **from univariates**

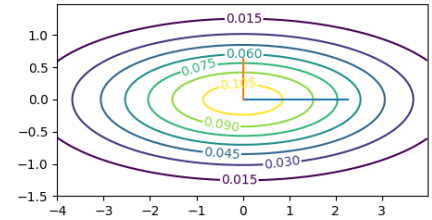
given n univariate Gaussians

$$\mathbf{X} \sim \mathcal{N}(0, I)$$

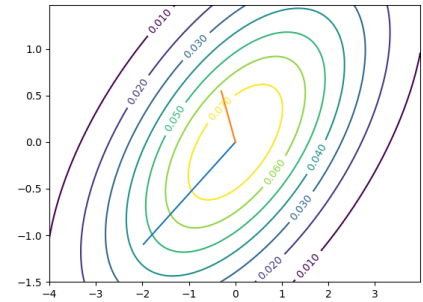


scale them by $\sqrt{D_{ii}}$

$$D^{\frac{1}{2}}\mathbf{X} \sim \mathcal{N}(0, D)$$

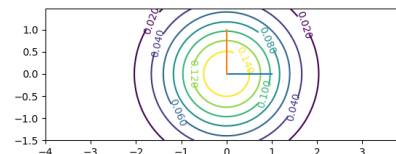


rotate/reflect using Q $QD^{\frac{1}{2}}\mathbf{X} \sim \mathcal{N}(0, QDQ^T) = \mathcal{N}(0, \Sigma)$

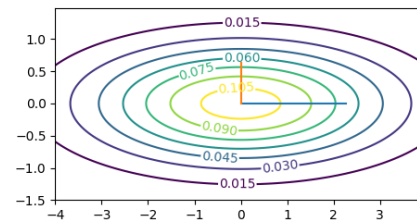


Multivariate Gaussian; from univariates

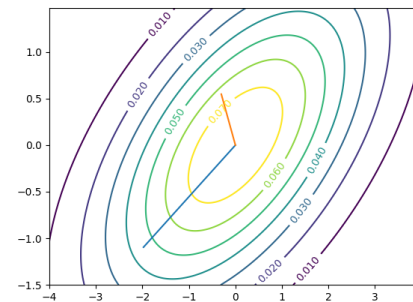
given n univariate Gaussians $\mathbf{X} \sim \mathcal{N}(0, I)$



scale them by $\sqrt{D_{ii}}$ $D^{\frac{1}{2}}\mathbf{X} \sim \mathcal{N}(0, D)$



rotate/reflect using Q $QD^{\frac{1}{2}}\mathbf{X} \sim \mathcal{N}(0, QDQ^T) = \mathcal{N}(0, \Sigma)$



more generally

$$\mathbf{X} \sim \mathcal{N}(\mu, \Sigma) \Rightarrow A\mathbf{X} + b \sim \mathcal{N}(A\mu + b, A\Sigma A^T)$$

parametrization

moment form (*mean parametrization*)

$$p(\mathbf{x}; \mu, \Sigma) = \frac{1}{\sqrt{|2\pi\Sigma|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mu)^T \Sigma^{-1}(\mathbf{x} - \mu)\right)$$

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↓
 $\eta = \Sigma^{-1}\mu$: local potential
 $\Lambda = \Sigma^{-1}$: precision matrix

parametrization

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↓

$$\eta = \Sigma^{-1} \mu : \text{local potential}$$
$$\Lambda = \Sigma^{-1} : \text{precision matrix}$$

information form (*cannonical parametrization*)

$$p(\mathbf{x}; \eta, \Lambda) = \sqrt{\frac{|\Lambda|}{(2\pi)^n}} \exp\left(-\frac{1}{2}\mathbf{x}^T \Lambda \mathbf{x} + \eta \mathbf{x} - \frac{1}{2}\eta^T \Lambda^{-1} \eta\right)$$

parametrization

moment form (*mean parametrization*)

$$p(\mathbf{x}; \mu, \Sigma) = \frac{1}{\sqrt{|2\pi\Sigma|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mu)^T \Sigma^{-1}(\mathbf{x} - \mu)\right)$$

↓

$$\begin{aligned} \eta &= \Sigma^{-1} \mu : \text{local potential} \\ \Lambda &= \Sigma^{-1} : \text{precision matrix} \end{aligned}$$

↑

$$\begin{aligned} \mu &= \Lambda^{-1} \eta \\ \Sigma &= \Lambda^{-1} \end{aligned}$$

information form (*cannonical parametrization*)

$$p(\mathbf{x}; \eta, \Lambda) = \sqrt{\frac{|\Lambda|}{(2\pi)^n}} \exp\left(-\frac{1}{2}\mathbf{x}^T \Lambda \mathbf{x} + \eta \mathbf{x} - \frac{1}{2}\eta^T \Lambda \eta\right)$$

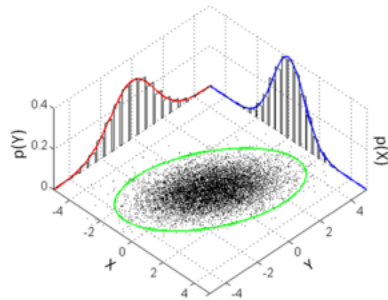
the relationship between the two types goes beyond Gaussians

(exp-family lecture)

marginalization

moment form $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

is useful for marginalization:



$$\mathbf{X} = [\mathbf{X}_A, \mathbf{X}_B]^T$$

$$\mathbf{X}_A \sim \mathcal{N}(\boldsymbol{\mu}_m, \boldsymbol{\Sigma}_m)$$

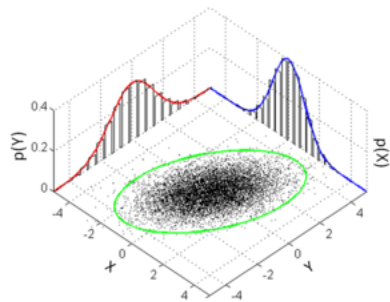
$$\boldsymbol{\mu} = [\boldsymbol{\mu}_A, \boldsymbol{\mu}_B]^T$$

$$\boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{AA} & \boldsymbol{\Sigma}_{AB} \\ \boldsymbol{\Sigma}_{BA} & \boldsymbol{\Sigma}_{BB} \end{bmatrix}$$

marginalization

moment form $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

is useful for marginalization:



$$\mathbf{X} = [\mathbf{X}_A, \mathbf{X}_B]^T$$

$$\mathbf{X}_A \sim \mathcal{N}(\boldsymbol{\mu}_m, \boldsymbol{\Sigma}_m)$$

$$\boldsymbol{\mu}_m = \boldsymbol{\mu}_A$$

$$\boldsymbol{\Sigma}_m = \boldsymbol{\Sigma}_A$$

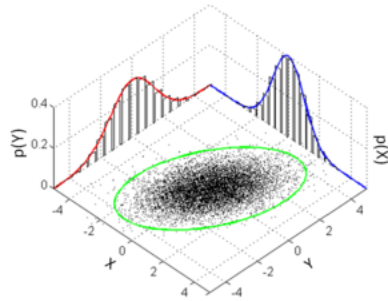
$$\boldsymbol{\mu} = [\boldsymbol{\mu}_A, \boldsymbol{\mu}_B]^T$$

$$\boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{AA} & \boldsymbol{\Sigma}_{AB} \\ \boldsymbol{\Sigma}_{BA} & \boldsymbol{\Sigma}_{BB} \end{bmatrix}$$

marginalization

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$$\mathbf{X} = [\mathbf{X}_A, \mathbf{X}_B]^T$$

$$\mathbf{X}_A \sim \mathcal{N}(\boldsymbol{\mu}_m, \boldsymbol{\Sigma}_m)$$

$$\boldsymbol{\mu}_m = \boldsymbol{\mu}_A$$

$$\boldsymbol{\Sigma}_m = \boldsymbol{\Sigma}_A$$

$$\boldsymbol{\mu} = [\boldsymbol{\mu}_A, \boldsymbol{\mu}_B]^T$$
$$\boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{AA} & \boldsymbol{\Sigma}_{AB} \\ \boldsymbol{\Sigma}_{BA} & \boldsymbol{\Sigma}_{BB} \end{bmatrix}$$

marginalization as a **linear transformation**: $A = [I_{AA}, 0]$

$$\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \Rightarrow A\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}_A, \boldsymbol{\Sigma}_{AA})$$

Marginal independencies; moment form

covariance means dependence & vice versa

$$X_i \perp X_j \mid \emptyset \Leftrightarrow \Sigma_{i,j} = \text{Cov}(X_i, X_j) = 0$$

why?

marginalize $\mathcal{N}(\mu, \Sigma)$ to get $\begin{bmatrix} X_i \\ X_j \end{bmatrix} \sim \mathcal{N}\left(\begin{bmatrix} \mu_i \\ \mu_j \end{bmatrix}, \begin{bmatrix} \sigma_i^2, \mathbf{0} \\ \mathbf{0}, \sigma_j^2 \end{bmatrix}\right) = \mathcal{N}(x_i; \mu_i, \sigma_i^2) \mathcal{N}(x_j; \mu_j, \sigma_j^2)$

Marginal independencies; moment form

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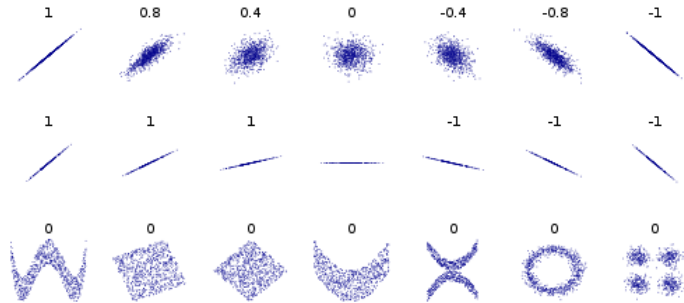


image from wikipedia

Gaussian is special in this sense

correlation: normalized covariance

$$\rho(X_i, X_j) = \frac{\text{Cov}(X_i, X_j)}{\sqrt{\text{Var}(X_i)\text{Var}(X_j)}}$$

Conditional independencies; *information form*

zeros of the **precision matrix** mean conditional independence

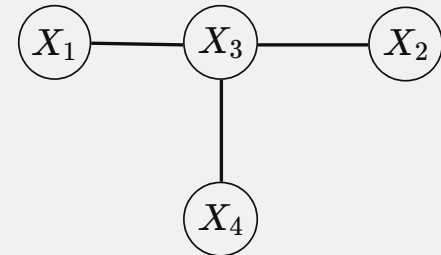
$$X_i \perp X_j \mid \mathbf{X} - \{X_i, X_j\} \Leftrightarrow \Lambda_{i,j} = 0$$

$\Lambda \neq 0$ adjacency matrix in the *Markov network* (**Gaussian MRF**)

$$p(\mathbf{x}; \eta, \Lambda) = \sqrt{\frac{|\Lambda|}{(2\pi)^n}} \exp\left(-\frac{1}{2}\mathbf{x}^T \Lambda \mathbf{x} + \eta \mathbf{x} - \frac{1}{2}\eta^T \Lambda \eta\right)$$

why?

$$\Lambda = \begin{bmatrix} \Lambda_{11}, & 0, & \Lambda_{1,3}, & 0 \\ 0, & \Lambda_{2,2}, & \Lambda_{2,3}, & 0 \\ \Lambda_{3,1}, & \Lambda_{3,2}, & \Lambda_{3,3}, & \Lambda_{3,4} \\ 0, & 0, & \Lambda_{4,3}, & \Lambda_{4,4} \end{bmatrix}$$



Conditional independencies; information form

zeros of the **precision matrix** mean conditional independence

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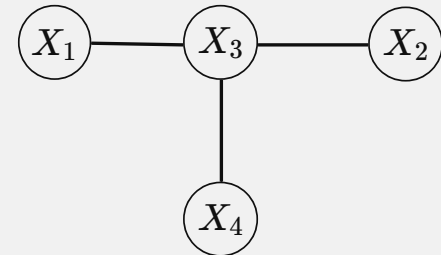
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why? write it as the product of factors:

$$\begin{array}{l} \text{corresponding} \\ \text{potentials} \end{array} \left| \begin{array}{l} \psi_{i,j}(x_i, x_j) = -x_i \Lambda_{i,j} x_j \\ \psi_i(x_i) = -\frac{1}{2} \Lambda_{i,i} x_i^2 + \eta_i x_i \end{array} \right.$$

$$\Lambda = \begin{bmatrix} \Lambda_{11} & 0 & \Lambda_{1,3} & 0 \\ 0 & \Lambda_{2,2} & \Lambda_{2,3} & 0 \\ \Lambda_{3,1} & \Lambda_{3,2} & \Lambda_{3,3} & \Lambda_{3,4} \\ 0 & 0 & \Lambda_{4,3} & \Lambda_{4,4} \end{bmatrix}$$



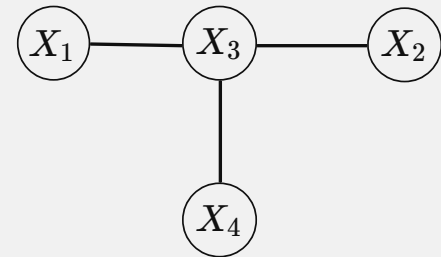
Gaussian MRF; information form

$$p(\mathbf{x}; \eta, \Lambda) = \sqrt{\frac{|\Lambda|}{(2\pi)^n}} \exp\left(-\frac{1}{2}\mathbf{x}^T \Lambda \mathbf{x} + \eta \mathbf{x} - \frac{1}{2}\eta^T \Lambda^{-1} \eta\right)$$

corresponding
potentials

$$\begin{cases} \psi_{i,j}(X_i, X_j) = -x_i \Lambda_{i,j} x_j \\ \psi_i(X_i) = -\Lambda_{i,i} x_i^2 + \eta_i x_i \end{cases}$$

$$\Lambda = \begin{bmatrix} \Lambda_{11}, & 0, & \Lambda_{1,3}, & 0 \\ 0, & \Lambda_{2,2}, & \Lambda_{2,3}, & 0 \\ \Lambda_{3,1}, & \Lambda_{3,2}, & \Lambda_{3,3}, & \Lambda_{3,4} \\ 0, & 0, & \Lambda_{4,3}, & \Lambda_{4,4} \end{bmatrix}$$



Λ should be **positive definite**

- otherwise the partition function

$$Z = \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2}\mathbf{x}^T \Lambda \mathbf{x} + \eta^T \mathbf{x}\right) d\mathbf{x} \text{ is not well-defined}$$

Conditioning; information form

marginalization: easy in the moment form

conditioning: easy in the information form

$$\mathbf{X}_A \mid \mathbf{X}_B \sim \mathcal{N}(\eta_{A|B}, \Lambda_{A|B})$$

why?

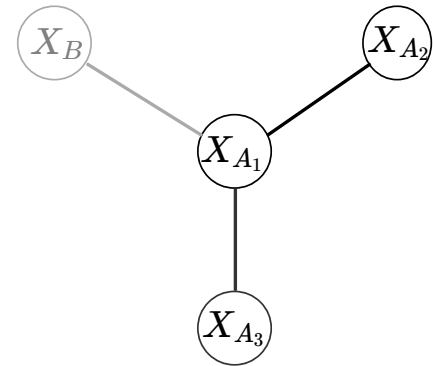
$$\Lambda_{A|B} = \Lambda_{AA}$$

$$\eta_{A|B} = \eta_A + \Lambda_{AB} \mathbf{X}_B$$

$$\mathbf{X} = [\mathbf{X}_A, \mathbf{X}_B]^T$$

$$\eta = [\eta_A, \eta_B]^T$$

$$\Lambda = \begin{bmatrix} \Lambda_{AA} & \Lambda_{AB} \\ \Lambda_{BA} & \Lambda_{BB} \end{bmatrix}$$



Conditioning; information form

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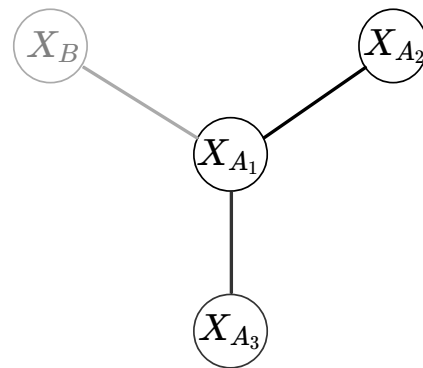
$$\Lambda = \begin{bmatrix} \Lambda_{AA} & \Lambda_{AB} \\ \Lambda_{BA} & \Lambda_{BB} \end{bmatrix}$$

not so easy in the moment form!

$$\mathbf{X}_A \mid \mathbf{X}_B \sim \mathcal{N}(\mu_{A|B}, \Sigma_{A|B})$$

$$\Sigma_{A|B} = \Sigma_{AA} - \Sigma_{AB} \Sigma_{BB}^{-1} \Sigma_{BA}$$

$$\mu_{A|B} = \mu_A - \Sigma_{AB} \Sigma_{BB}^{-1} (\mathbf{X}_B - \mu_B)$$

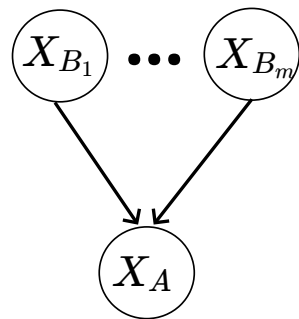


Gaussian Bayesian network

an alternative representation for multivariate Gaussian

linear Gaussian CPD $X_A | \mathbf{X}_B \sim \mathbf{w}^T \mathbf{X}_B + \mathcal{N}(\mu_0, \sigma_0^2)$

and $\mathbf{X}_B \sim \mathcal{N}(\mu_B, \Sigma_B)$



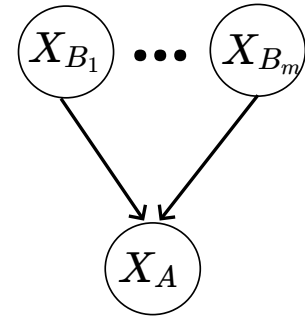
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joint dist. \Leftrightarrow conditional form (CPD)



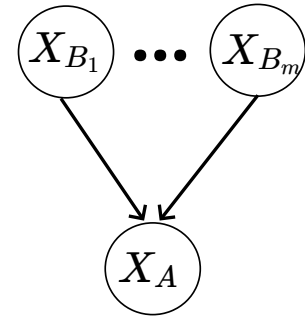
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joint dist. \Leftrightarrow conditional form (CPD)



$$X_0 \sim \mathcal{N}(\mu_0, \sigma_0^2)$$

$$X_1 = \mathbf{w}^T \mathbf{X}_B \sim \mathcal{N}(\mathbf{w}^T \mu_B, \mathbf{w}^T \Sigma_B \mathbf{w})$$

$$X_A = X_0 + X_1 \sim \mathcal{N}(\mu_0 + \mathbf{w}^T \mu_B, \sigma_0^2 + \mathbf{w}^T \Sigma_B \mathbf{w})$$

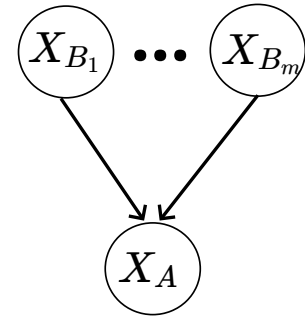
Gaussian Bayesian network

an alternative representation for multivariate Gaussian

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and $\mathbf{X}_B \sim \mathcal{N}(\mu_B, \Sigma_B)$

joint dist. \Leftrightarrow conditional form (CPD)



sum of two Gaussian RVs is a Gaussian RV

the pdf of the sum of RVs from the convolution of pdfs

$$X_0 \sim \mathcal{N}(\mu_0, \sigma_0^2)$$

$$X_1 = \mathbf{w}^T \mathbf{X}_B \sim \mathcal{N}(\mathbf{w}^T \mu_B, \mathbf{w}^T \Sigma_B \mathbf{w})$$

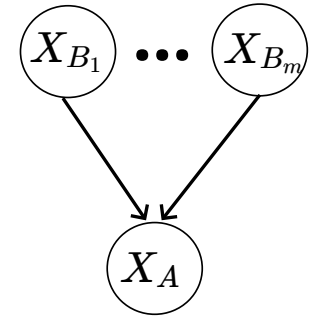
$$X_A = X_0 + X_1 \sim \mathcal{N}(\mu_0 + \mathbf{w}^T \mu_B, \sigma_0^2 + \mathbf{w}^T \Sigma_B \mathbf{w})$$

Gaussian Bayesian network

an alternative representation for multivariate Gaussian

linear Gaussian CPD $X_A | \mathbf{X}_B \sim \mathbf{w}^T \mathbf{X}_B + \mathcal{N}(\mu_0, \sigma_0^2)$

and $\mathbf{X}_B \sim \mathcal{N}(\mu_B, \Sigma_B)$



joint dist. \Leftrightarrow conditional form (CPD)

marginal over X_A is $X_A \sim \mathcal{N}(\mu_0 + \mathbf{w}^T \mu_B, \sigma_0^2 + \mathbf{w}^T \Sigma_B \mathbf{w})$

joint dist. is $(X_A, \mathbf{X}_B) \sim \mathcal{N} \left(\begin{bmatrix} \mu_0 + \mathbf{w}^T \mu_B \\ \mu_B \end{bmatrix}, \begin{bmatrix} \sigma_0^2 + \mathbf{w}^T \Sigma_B \mathbf{w}, & \mathbf{w}^T \Sigma_B \\ \Sigma_B \mathbf{w}, & \Sigma_B \end{bmatrix} \right)$

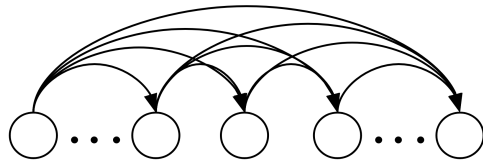
all the other elements follow from the marginals

$$\text{Cov}(X_A, X_{B,i}) = \sum_j w_j \text{Cov}(X_{B,j}, X_{B,i})$$

Gaussian Bayesian network

an alternative representation for multivariate Gaussian

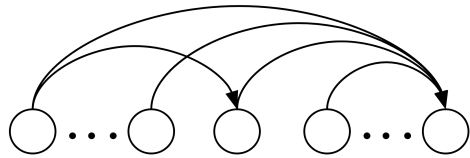
linear Gaussian CPD $X_i | Pa_{X_i} \sim \mathbf{w}_i^T Pa_{X_i} + \mathcal{N}(\mu_i, \sigma_i^2)$



worst case: $\mathcal{O}(n)$ parameters per node

- even if Σ is sparse!

generally:

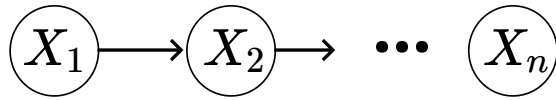


- DAG structure depends on the ordering
- v-structures $X_k \perp X_j \Rightarrow \Sigma_{k,j} = 0$
- d-separation to find the sparsity of Σ

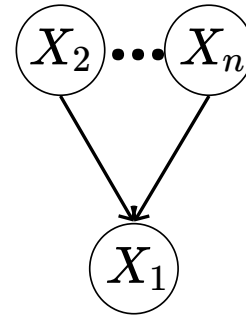
quiz

what is the sparsity patterns of Σ , Λ , and \mathbf{w}^i in Gaussian BN?

case 1



case 2



Summary

- multivariate Gaussian:
 - mean param. (moment form) Σ, μ
 - useful for marginalization
 - sparsity \Leftrightarrow *marginal* independence
 - canonical param. (information form) Λ, η
 - useful for conditioning
 - sparsity \Leftrightarrow *conditional* independence
- Gaussian Bayesian network (linear Gaussian CPD)
- Gaussian MRF (using information form)